# Control randomisation approach for policy gradient and application to reinforcement learning in optimal switching 

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#### Abstract

We propose a comprehensive framework for policy gradient methods tailored to continuous time reinforcement learning. This is based on the connection between stochastic control problems and randomised problems, enabling applications across various classes of Markovian continuous time control problems, beyond diffusion models, including e.g. regular, impulse and optimal stopping/switching problems. By utilizing change of measure in the control randomisation technique, we derive a new policy gradient representation for these randomised problems, featuring parametrised intensity policies. We further develop actor-critic algorithms specifically designed to address general Markovian stochastic control issues. Our framework is demonstrated through its application to optimal switching problems, with two numerical case studies in the energy sector focusing on real options.


Key words: Reinforcement learning in continuous time, policy gradient, control randomization, actor-critic algorithms, optimal switching,

[^0]
## 1 Introduction

The theory of reinforcement learning for continuous time stochastic control has advanced significantly, beginning with the foundational work [17], and continuing with [11, [10] who developed policy gradient methods and actor-critic algorithms, and [12] for $q$-learning. These studies primarily focus on regular controls within diffusion processes, employing the Feynman-Kac formula and the partial differential equations (PDE) representation of the value function to derive gradients of the performance value function with respect to the parameters of the stochastic policy.

Our research aims to expand the application of these methods beyond diffusion models to a broader range of Markovian control problems, including singular, impulse, and optimal stopping and switching problems. To achieve this, we propose a unified framework with a general reformulation in terms of Markovian randomised problems. This approach to stochastic control problem is commonly referred to as control randomisation, initially introduced in [1] for optimal switching problems, and further developed in [13] for impulse control, in [14] for regular controls, and in (5) for general non-Markovian stochastic control problems. The basic idea is to replace the control process $\left(\alpha_{t}\right)_{t}$ valued in $A$ by a random (uncontrolled) point process $\left(I_{t}\right)_{t}$ with marks in $A$, formulate an auxiliary control problem where the intensity distribution of $I$ is controlled, called randomized problem, and show that the value functions of the two problems coincide. The key feature of the randomised problems is its formulation in terms of a family of dominated probability measures under which the optimization is performed.

Utilizing the change of measure in these randomised settings, we derive a gradient representation of the value function with respect to parametrised intensity policies directly, without reliance on PDEs. This framework not only incorporates Poisson discretisation as per the randomization method but also accommodates standard fixed discretisations for continuous-time problems. Using this policy gradient, we design an Actor-Critic algorithm to alternately learn the value function and the optimal intensity policy. Notably, the gradient structure relies solely on the state at action points, circumventing the need for further discretisation during implementation.

We demonstrate the applicability of our results in a model-free setting, learning optimal control and value functions through empirical observations and samples. This methodology is applied specifically to optimal switching problems but is adaptable to a wide variety of continuous stochastic control scenarios. We provide numerical examples from real options in the energy markets to illustrate these concepts.

The remainder of the paper is structured as follows: In Section 2, we detail the Markovian randomised problem and develop a corresponding policy gradient method. Section 3 applies this methodology to a diverse array of continuous time control problems, and Section 4 presents and evaluates numerical experiments within the context of optimal switching problems.

## 2 Policy gradient method for Markovian randomised problems

In this section, we will consider a general class of Markovian randomised control problems in continuous time. Control randomisation method can be seen as a unified approach to a large class of control problems in continuous time, including optimal stopping, switching and impulse control problems, as we will see in Section 3. We will derive first a general policy gradient representation and then from this, an Actor-Critic algorithm to tackle this class of problems.

### 2.1 Theory background

### 2.1.1 Randomised control problem setup

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space on which we consider a simple random counting measure $\nu$ on $(0, \infty) \times A$ with $A$ some Polish space, such that $\mathbb{E}[\nu((0, T] \times A)]<\infty$, and associated to the marked point process $\left(\tau_{n}, \mathrm{a}_{n}\right)_{n}$ and the pure jump $A$-valued process $I$ with dynamics

$$
\begin{equation*}
d I_{s}=\int_{A}\left(e-I_{s-}\right) \nu(d s, d e), \quad s \leq T \tag{2.1}
\end{equation*}
$$

We denote by $I^{t, a}$ the jump process starting from $a \in A$, at time $t \in[0, T]$, that is $I_{t}^{t, a}=a$, and following the dynamics (2.1) for $t \leq s \leq T$. We consider a state process $X$ valued on $\mathbb{R}^{d}$ s.t. the pair ( $X, I$ ) is Markov, and $X$ only jumps at the times given by $\nu$. An example includes the case where $X$ is driven by a SDE in the form

$$
\begin{equation*}
d X_{s}=b\left(s, X_{s}, I_{s}\right) d s+\sigma\left(s, X_{s}, I_{s}\right) d W_{s}+\int_{A} \gamma\left(s, X_{s-}, I_{s-}, e\right) \nu(d s, d e) \tag{2.2}
\end{equation*}
$$

with $W$ a Brownian motion. We denote by $X^{t, x, a}$ the state process $X$ that starts from $x$ at time $t$, that is $X_{t}^{t, x, a}=x$, and s.t. $\left(X^{t, x, a}, I^{t, a}\right)$ is Markov, and we assume the estimate

$$
\mathbb{E}\left[\sup _{t \leq s \leq T}\left|X_{s}^{t, x, a}\right|^{p}\right] \leq C\left(1+|x|^{p}\right), \quad \forall x \in \mathbb{R}^{d}
$$

for some positive constant $C$ and $p \in[1, \infty)$. This estimate is satisfied for $X$ as in (2.2) under standard Lipschitz and linear growth conditions on $b, \sigma, \gamma$.

By [9, Theorems 2.1, 2.3, 3.4], there exists a unique (up to a $\mathbb{P}$-null set) predictable random measure $\hat{\nu}$ with $\hat{\nu}(\{s\} \times A) \leq 1$ for all $s \in(0, T]$ such that for every $\mathcal{P}\left(\mathbb{F}^{X, \nu}\right) \otimes \mathcal{B}(A)$-measurable random field $H \geq 0$, where $\mathcal{P}\left(\mathbb{F}^{X, \nu}\right)$ denotes the predictable $\sigma$-algebra of $\mathbb{F}^{X, \nu}$, it holds that

$$
\mathbb{E}\left[\int_{0}^{T} \int_{A} H(s, e) \nu(d s, d e)\right]=\mathbb{E}\left[\int_{0}^{T} \int_{A} H(s, e) \hat{\nu}(d s, d e)\right],
$$

called the predictable projection or compensator of $\nu$, which is uniquely characterising $\nu$. Guided by the approach of the randomisation method, we will now optimise over the set of (in a suitable sense) "intensities" of the process $I$. To this end, we note that for every $\mathbb{F}^{X, \nu} \otimes \mathcal{B}(A)$-predictable, essentially bounded process $\lambda$ satisfying
(i) $\int_{A} \lambda_{s}(e) \hat{\nu}(\{s\}, d e) \leq 1$ for all $s \in(0, T]$,
(ii) for all $s \in(0, T]$ such that $\hat{\nu}(\{s\} \times A)=1$, it also holds that $\int_{A} \lambda_{s}(e) \hat{\nu}(\{s\}, d e)=1$, we can construct a tilted probability measure $\mathbb{P}^{\lambda} \ll \mathbb{P}$ such that $\nu$ is a random point measure with the predictable projection $\lambda_{s}(e) \hat{\nu}(d s, d e)$ under $\mathbb{P}^{\lambda}$. This is achieved through Girsanov's theorem, as outlined in e.g. [9, Theorem 4.5], by defining $\mathbb{P}^{\lambda}$ via its density process

$$
\begin{align*}
Z_{s}^{\lambda}:=\left.\frac{d \mathbb{P}^{\lambda}}{d \mathbb{P}^{\lambda}}\right|_{\mathcal{F}_{s}^{X, \nu}}= & \prod_{t \in(0, s], 0<\hat{\nu}(\{t\} \times A)<1, \nu(\{t\} \times A)=0} \frac{1-\int_{A} \lambda_{t}(e) \hat{\nu}(\{t\}, d e)}{1-\hat{\nu}(\{t\} \times A)} \\
& \cdot \exp \left(\int_{(0, s]} \int_{A} \log \lambda_{t}(e) \nu(d t, d e)-\int_{(0, s]} \int_{A}\left(\lambda_{t}(e)-1\right) \hat{\nu}^{c}(d t, d e)\right), \tag{2.3}
\end{align*}
$$

for $s \in(0, T]$, where $\hat{\nu}^{c}(d s, d e):=\mathbb{1}_{\{\hat{\nu}(\{s\} \times A)=0\}} \hat{\nu}(d s, d e)$.
Notice that we do not assume necessarily that the compensator is absolutely continuous w.r.t. the Lebesgue measure $d s$, in order to take into account the possibility of jumps at deterministic times, hence to embed the case of stochastic control on discrete time, i.e. Markov decision process.

By the Markovian structure of our problem, we now define the set of admissible control $\mathcal{V}$ as all such processes $\lambda$ satisfying the above conditions while being of the form

$$
\lambda_{s}(e)=\lambda\left(e \mid s, X_{s-}, I_{s-}\right), \quad s \leq T,
$$

for some bounded deterministic function $\lambda$ on $A \times[0, T] \times \mathbb{R}^{d} \times A$. For the ease of arguments, we furthermore require that all $\lambda \in \mathcal{V}$ satisfy the following conditions which ensure that also $\mathbb{P} \ll \mathbb{P}^{\lambda}$ and thus $\mathbb{P}^{\lambda} \sim \mathbb{P},{ }^{1}$
(i) $\lambda$ is bounded away from 0 , that is $\inf _{(s, x, a, e)} \lambda(e \mid s, x, a)>0$,
(ii) there exists a constant $C<1$ such that for all $s \in(0, T]$, when $\hat{\nu}(\{s\} \times A)<1$, then it also holds that $\int_{A} \lambda\left(e \mid s, X_{s-}, I_{s-}\right) \hat{\nu}(\{s\}, d e) \leq C<1$.

Note that for each such $\lambda \in \mathcal{V}$, the process $(X, I)$ will still be Markovian under $\mathbb{P}^{\lambda} \ll \mathbb{P}$, and we have the estimate

$$
\mathbb{E}^{\lambda}\left[\sup _{t \leq s \leq T}\left|X_{s}^{t, x, a}\right|^{p}\right] \leq C_{\lambda}\left(1+|x|^{p}\right), \quad \forall x \in \mathbb{R}^{d},
$$

where $\mathbb{E}^{\lambda}$ denotes the expectation under $\mathbb{P}^{\lambda}$. In general, our objective is now to optimise the reward functional:
$J(t, x, a, \lambda):=\mathbb{E}^{\lambda}\left[g\left(X_{T}^{t, x, a}, I_{T}^{t, a}\right)+\int_{t}^{T} f\left(s, X_{s}^{t, x, a}, I_{s}^{t, a}\right) d s-\int_{(t, T]} \int_{A} c\left(s, X_{s-}^{t, x, a}, I_{s-}^{t, a}, e\right) \nu(d s, d e)\right]$,
for $(t, x, a) \in[0, T] \times \mathbb{R}^{d} \times A$, where the reward functions $f, g$ and the cost function $c$ are assumed to satisfy the polynomial growth condition

$$
|f(t, x, a)|+|g(x, a)|+|c(t, x, a, e)| \leq C\left(1+|x|^{p}\right),
$$

for all $t \in[0, T], x \in \mathbb{R}^{d}, a, e \in A$. Notice that the reward functional $J$ then also satisfies the the polynomial growth condition

$$
|J(t, x, a, \lambda)| \leq C_{\lambda}\left(1+|x|^{p}\right) .
$$

Remark 2.1. From the definition of the reward functional $J$ and the Markov property of ( $X, I$ ), we have the martingale property under $\mathbb{P}^{\lambda}, \lambda \in \mathcal{V}$, of the process

$$
J\left(s, X_{s}, I_{s}, \lambda\right)+\int_{0}^{s} f\left(r, X_{r}, I_{r}\right) d r-\int_{(0, s]} \int_{A} c\left(r, X_{r-}, I_{r-}, e\right) \nu(d r, d e), \quad 0 \leq s \leq T .
$$

[^1]
### 2.1.2 Policy gradient representation

The policy gradient method aims to optimize the expected reward $J$ by exploring a parameterized family $\left(\lambda^{\theta}\right)_{\theta \in \Theta} \subseteq \mathcal{V}$. This family is chosen to be sufficiently dense in $\mathcal{V}$, meaning that

$$
\sup _{\lambda \in \mathcal{V}} J(t, x, a, \lambda)=\sup _{\theta \in \Theta} J(t, x, a, \theta)
$$

where we denote by $J(t, x, a, \theta):=J\left(t, x, a, \lambda^{\theta}\right)$ with a slight abuse of notation. The optimization process then involves computing the gradient of $J^{\theta}$ with respect to the parameter $\theta$, allowing for updates to the policy parameters - typically done through methods like gradient descent to maximize the overall reward.

Our aim in this section is now to derive an explicit formula for the gradient $\nabla_{\theta} J^{\theta}(t, x, a, \theta)$. While the approach by 11] is based on the Feynman-Kac formula for $J^{\theta}$, we will instead use the Girsanov formula (2.3). The advantage is that we do not need to assume or impose conditions for ensuring regularity on the functional $J$ for deriving the partial differential equations that it satisfies in the continuous-time framework. This is crucial since the function $J$ may be discontinuous in time in the case where $\hat{\nu}$ admits atoms in time, and then PDE method cannot be applied.

We shall assume that for all $(t, x, a, e) \in[0, T] \times \mathbb{R}^{d} \times A \times A$, the map $\theta \in \Theta \mapsto \lambda^{\theta}(e \mid t, x, a)$ is differentiable with a derivative satisfying the growth condition: for each $\theta \in \Theta$, there exists some positive constant $C_{\theta}$ s.t.

$$
\int_{(t, T]} \int_{A}\left|\nabla_{\theta} \lambda^{\theta}(e \mid s, x, a)\right| \hat{\nu}(d s, d e) \leq C_{\theta}(1+|x|), \quad(t, x) \in[0, T] \times \mathbb{R}^{d}
$$

Theorem 2.2. We have

$$
\begin{align*}
\nabla_{\theta} J(t, x, a, \theta)=\mathbb{E}^{\theta}[ & \int_{(t, T]} \int_{A} \nabla_{\theta}\left(\log \lambda^{\theta}\right)\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \\
& \left.\cdot\left(J\left(s, X_{s}^{t, x, a}, e, \theta\right)-J\left(s, X_{s-}^{t, x, a}, I_{s-}^{t, a}, \theta\right)-c\left(s, X_{s-}^{t, x, a}, I_{s-}^{t, a}, e\right)\right) \nu(d s, d e)\right] \tag{2.4}
\end{align*}
$$

for $(t, x, a) \in[0, T] \times \mathbb{R}^{d} \times A$, where $\mathbb{E}^{\theta}$ denotes the expectation under $\mathbb{P}^{\theta}=\mathbb{P}^{\lambda^{\theta}}$.
Proof. From Bayes formula with (2.3), the reward functional is formulated in term of the reference probability measure $\mathbb{P}$ instead of $\mathbb{P}^{\theta}:=\mathbb{P}^{\lambda^{\theta}}$ as follows for $\theta \in \Theta$,

$$
\begin{aligned}
& J(t, x, a, \theta) \\
= & \mathbb{E}^{\theta}\left[g\left(X_{T}^{t, x, a}, I_{T}^{t, a}\right)+\int_{t}^{T} f\left(s, X_{s}^{t, x, a}, I_{s}^{t, a}\right) d s-\int_{(t, T]} \int_{A} c\left(s, X_{s-}^{t, x, a}, I_{s-}^{t, a}, e\right) \nu(d s, d e)\right] \\
= & \mathbb{E}\left[Z_{T}^{t, x, a, \theta}\left(g\left(X_{T}^{t, x, a}, I_{T}^{t, a}\right)+\int_{t}^{T} f\left(s, X_{s}^{t, x, a}, I_{s}^{t, a}\right) d s-\int_{(t, T]} \int_{A} c\left(s, X_{s-}^{t, x, a}, I_{s-}^{t, a}, e\right) \nu(d s, d e)\right)\right]
\end{aligned}
$$

where

$$
\begin{aligned}
Z_{T}^{t, x, a, \theta}=\exp & \left(\int_{(t, T]} \int_{A} \log \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \nu(d s, d e)-\int_{(t, T]} \int_{A}\left(\lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right)-1\right) \hat{\nu}^{c}(d s, d e)\right. \\
& \left.+\sum_{s \in(t, T], 0<\hat{\nu}(\{s\} \times A)<1} \mathbb{1}_{\{\nu(\{s\} \times A)=0\}} \log \left(\frac{1-\int_{A} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)}{1-\hat{\nu}(\{s\} \times A)}\right)\right)
\end{aligned}
$$

By differentiating this relation w.r.t. $\theta$, and writing $\nabla_{\theta} Z_{T}^{t, x, a, \theta}=Z_{T}^{t, x, a, \theta} L_{T}^{t, x, a, \theta}$ with

$$
\begin{aligned}
L_{T}^{t, x, a, \theta}= & \int_{(t, T]} \int_{A} \nabla_{\theta}\left(\log \lambda^{\theta}\right)\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \nu(d s, d e)-\int_{(t, T]} \int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}^{c}(d s, d e) \\
& -\sum_{s \in(t, T], 0<\hat{\nu}(\{s\} \times A)<1} \mathbb{1}_{\{\nu(\{s\} \times A)=0\}} \frac{\int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)}{1-\int_{A} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)}
\end{aligned}
$$

we get

$$
\begin{aligned}
& \nabla_{\theta} J(t, x, a, \theta) \\
= & \mathbb{E}^{\theta}\left[L_{T}^{t, x, a, \theta}\left(g\left(X_{T}^{t, x, a}, I_{T}^{t, a}\right)+\int_{t}^{T} f\left(s, X_{s}^{t, x, a}, I_{s}^{t, a}\right) d s-\int_{(t, T]} \int_{A} c\left(s, X_{s-}^{t, x, a}, I_{s-}^{t, a}, e\right) \nu(d s, d e)\right)\right]
\end{aligned}
$$

To simplify this expression, we will use that due the Markovian structure of our problem, the process $M^{t, x, a, \theta}$ given by

$$
M_{s}^{t, x, a, \theta}:=J\left(s, X_{s}^{t, x, a}, I_{s}^{t, a}, \theta\right)+\int_{t}^{s} f\left(r, X_{r}^{t, x, a}, I_{r}^{t, a}\right) d r-\int_{(t, s]} \int_{A} c\left(r, X_{r-}^{t, x, a}, I_{r-}^{t, a}, e\right) \nu(d r, d e), \quad s \in[t, T]
$$

is a $\mathbb{P}^{\theta}$-martingale, see also Remark 2.1. We start by noting that $J\left(T, X_{T}^{t, x, a}, I_{T}^{t, a}, \theta\right)=g\left(X_{T}^{t, x, a}, I_{T}^{t, a}\right)$, which allows us to write $\nabla_{\theta} J(t, x, a, \theta)$ using $M^{t, x, a, \theta}$ as follows

$$
\begin{aligned}
\nabla_{\theta} J(t, x, a, \theta)= & \mathbb{E}^{\theta}\left[L_{T}^{t, x, a, \theta} M_{T}^{t, x, a, \theta}\right] \\
=\mathbb{E}^{\theta} & {\left[\int_{(t, T]} \int_{A} \nabla_{\theta}\left(\log \lambda^{\theta}\right)\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) M_{T}^{t, x, a, \theta} \nu(d s, d e)\right.} \\
& -\int_{(t, T]} \int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) M_{T}^{t, x, a, \theta} \hat{\nu}^{c}(d s, d e) \\
& \left.-\sum_{s \in(t, T], 0<\hat{\nu}(\{s\} \times A)<1} \mathbb{1}_{\{\nu(\{s\} \times A)=0\}} \frac{\int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)}{1-\int_{A} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)} M_{T}^{t, x, a, \theta}\right] .
\end{aligned}
$$

Now using the $\mathbb{P}^{\theta}$-martingale property of $M^{t, x, a, \theta}$, we obtain

$$
\begin{align*}
\nabla_{\theta} J(t, x, a, \theta)=\mathbb{E}^{\theta} & {\left[\int_{(t, T]} \int_{A} \nabla_{\theta}\left(\log \lambda^{\theta}\right)\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) M_{s}^{t, x, a, \theta} \nu(d s, d e)\right.} \\
& -\int_{(t, T]} \int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) M_{s}^{t, x, a, \theta} \hat{\nu}^{c}(d s, d e) \\
& \left.-\sum_{s \in(t, T], 0<\hat{\nu}(\{s\} \times A)<1} \mathbb{1}_{\{\nu(\{s\} \times A)=0\}} \frac{\int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)}{1-\int_{A} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)} M_{s}^{t, x, a, \theta}\right] \tag{2.5}
\end{align*}
$$

To simplify the notation in the following arguments, let us introduce the predictable process
$N_{s}^{t, x, a, \theta}:=J\left(s, X_{s-}^{t, x, a}, I_{s-}^{t, a}, \theta\right)+\int_{t}^{s} f\left(r, X_{r}^{t, x, a}, I_{r}^{t, a}\right) d r-\int_{(t, s)} \int_{A} c\left(r, X_{r-}^{t, x, a}, I_{r-}^{t, a}, e\right) \nu(d r, d e), \quad s \in[t, T]$.
We note that $\nu$ and thus also $I^{t, a}$ and $X^{t, x, a}$ have $\mathbb{P}$-a.s. only finitely many discontinuities on $[t, T]$. This implies that $\left\{s \in[t, T] \mid M_{s}^{t, x, a, \theta} \neq N_{s}^{t, x, a, \theta}\right\}$ is $\mathbb{P}$-a.s. countable and thus $M_{s}^{t, x, a, \theta}=N_{s}^{t, x, a, \theta}$, $\mathbb{P} \otimes \hat{\nu}^{c}(\cdot, A)$-a.s., which allows us to rewrite the second term in (2.5) as

$$
\begin{align*}
& \mathbb{E}^{\theta}\left[\int_{(t, T]} \int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) M_{s}^{t, x, a, \theta, \theta} \hat{\nu}^{c}(d s, d e)\right] \\
& =\mathbb{E}^{\theta}\left[\int_{(t, T]} \int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) N_{s}^{t, x, a, \theta_{\hat{\nu}}} \hat{\nu}^{c}(d s, d e)\right] . \tag{2.6}
\end{align*}
$$

For the last term in (2.5), focusing on the not-almost-sure jumps of $\nu$, and using that $\nu(\{s\} \times A)=$ 0 implies that $I_{s-}^{t, a}=I_{s}^{t, a}$ and then by assumption also $X_{s-}^{t, x, a}=X_{s}^{t, x, a}$, which implies that $M_{s}^{t, x, a, \theta}=N_{s}^{t, x, a, \theta}$, we obtain

$$
\begin{align*}
& \mathbb{E}^{\theta}\left[\sum_{s \in(t, T], 0<\hat{\nu}(\{s\} \times A)<1} \mathbb{1}_{\{\nu(\{s\} \times A)=0\}} \frac{\int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s}^{t, a}\right) \hat{\nu}(\{s\}, d e)}{1-\int_{A} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)} M_{s}^{t, x, a, \theta}\right] \\
& =\mathbb{E}^{\theta}\left[\sum_{s \in(t, T], 0<\hat{\nu}(\{s\} \times A)<1}\left(1-\mathbb{1}_{\{\nu(\{s\} \times A)>0\}}\right) \frac{\int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)}{1-\int_{A} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)} N_{s}^{t, x, a, \theta}\right] . \tag{2.7}
\end{align*}
$$

To simply this term, we note that since $\nu$ is a simple random counting measure

$$
\begin{aligned}
& \mathbb{E}^{\theta}\left[\sum_{s \in(t, T], 0<\hat{\nu}(\{s\} \times A)<1} \mathbb{1}_{\{\nu(\{s\} \times A)>0\}} \frac{\int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s}^{t, x, a}, I_{s}^{t, a}\right) \hat{\nu}(\{s\}, d e)}{1-\int_{A} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)} N_{s}^{t, x, a, \theta}\right] \\
& =\mathbb{E}^{\theta}\left[\int_{(t, T]} \int_{A} \mathbb{1}_{\{0<\hat{\nu}(\{s\} \times A)<1\}} \frac{\int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s}^{t, a}\right) \hat{\nu}(\{s\}, d e)}{1-\int_{A} \lambda^{\theta}\left(e \mid s, X_{s}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)} N_{s}^{t, x, a, \theta} \nu(\{s\}, d u)\right] \\
& =\mathbb{E}^{\theta}\left[\sum_{s \in(t, T], 0<\hat{\nu}(\{s\} \times A)<1} \frac{\int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)}{1-\int_{A} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)} N_{s}^{t, x, a, \theta} \int_{A} \lambda^{\theta}\left(u \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d u)\right],
\end{aligned}
$$

using that $N^{t, x, a, \theta}$ is by construction predictable. Therefore, we can rewrite (2.7) as

$$
\begin{align*}
& \mathbb{E}^{\theta}\left[\sum_{s \in(t, T], 0<\hat{\nu}(\{s\} \times A)<1} \mathbb{1}_{\{\nu(\{s\} \times A)=0\}} \frac{\int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s}^{t, a}\right) \hat{\nu}(\{s\}, d e)}{1-\int_{A} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)} M_{s}^{t, x, a, \theta}\right] \\
& =\mathbb{E}^{\theta}\left[\sum_{s \in(t, T], 0<\hat{\nu}(\{s\} \times A)<1} \int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) N_{s}^{t, x, a, \theta, \theta} \hat{\nu}(\{s\}, d e)\right] . \tag{2.8}
\end{align*}
$$

To continue, we need an auxiliary result, for which we will take a closer look at the times where $\hat{\nu}(\{s\} \times A)=1$. Since $\lambda^{\theta}$ is an admissible control, this implies for such time points that
$\int_{A} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-a}^{t, a}\right) \hat{\nu}(\{s\}, d e)=1$ for all $\theta \in \Theta$ and thus

$$
\int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)=\nabla_{\theta}\left(\int_{A} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)\right)=0 .
$$

This leads us to

$$
\begin{equation*}
\mathbb{E}^{\theta}\left[\sum_{s \in(t, T], \hat{\nu}(\{s\} \times A)=1} \int_{A} N_{s}^{t, x, a, \theta} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)\right]=0 . \tag{2.9}
\end{equation*}
$$

Thus, putting (2.6), (2.8) and (2.9) together, we obtain that

$$
\begin{aligned}
\mathbb{E}^{\theta} & {\left[\int_{(t, T]} \int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) M_{s}^{t, x, a, \theta} \hat{\nu}^{c}(d s, d e)\right.} \\
& \left.+\sum_{s \in(t, T], 0<\hat{\nu}(\{s\} \times A)<1} \mathbb{1}_{\{\nu(\{s\} \times A)=0\}} \frac{\int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)}{1-\int_{A} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(\{s\}, d e)} M_{s}^{t, x, a, \theta}\right] \\
=\mathbb{E}^{\theta} & {\left[\int_{(t, T]} \int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) N_{s}^{t, x, a, \theta} \hat{\nu}(d s, d e)\right] . }
\end{aligned}
$$

Now using that the integrand $N^{t, x, a, \theta}$ is predictable, we can apply that $\lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \hat{\nu}(d s, d e)$ is the predictable projection of $\nu$ under $\mathbb{P}^{\theta}$, to obtain

$$
\begin{aligned}
& \mathbb{E}^{\theta}\left[\int_{(t, T]} \int_{A} \nabla_{\theta} \lambda^{\theta}\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) N_{s}^{t, x, a, \theta} \hat{\nu}(d s, d e)\right] \\
& =\mathbb{E}^{\theta}\left[\int_{(t, T]} \int_{A} \nabla_{\theta}\left(\log \lambda^{\theta}\right)\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) N_{s}^{t, x, a, \theta} \nu(d s, d e)\right] .
\end{aligned}
$$

Finally, together with (2.5), we obtain

$$
\begin{aligned}
\nabla_{\theta} J(t, x, a, \theta)=\mathbb{E}^{\theta} & {\left[\int_{(t, T]} \int_{A} \nabla_{\theta}\left(\log \lambda^{\theta}\right)\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right) \cdot\left(M_{s}^{t, x, a, \theta}-N_{s}^{t, x, a, \theta}\right) \nu(d s, d e)\right] } \\
=\mathbb{E}^{\theta} & {\left[\int_{(t, T]} \int_{A} \nabla_{\theta}\left(\log \lambda^{\theta}\right)\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right)\right.} \\
& \left.\cdot\left(J\left(s, X_{s}^{t, x, a}, e, \theta\right)-J\left(s, X_{s-}^{t, x, a}, I_{s-}^{t, a}, \theta\right)-c\left(s, X_{s-}^{t, x, a}, I_{s-,}^{t, a}, e\right)\right) \nu(d s, d e)\right] .
\end{aligned}
$$

We will see that we can use this to construct policy gradient (PG) steps for a diverse class of control problems in continuous time.

### 2.2 Actor-critic algorithm

We now aim to design an actor-critic (AC) learning algorithm for our randomised problem. AC algorithms are useful to tackle problems in environments where explicit knowledge of system dynamics is unavailable (e.g. model-free settings) and they consist out of two steps which are executed in turns: the policy evaluation (PE) step updates our reward functional estimate $J$ based on the current policy $\lambda$, and the policy gradient (PG) step updates our current policy $\lambda$
using the current estimate of $J$. This enables simultaneous learning of the optimal parameters $\kappa$ and $\theta$ for our parametrised families $\left(J^{\kappa}\right)_{\kappa}$ representing the reward functional and $\left(\lambda^{\theta}\right)_{\theta}$ for the optimal intensity control. In particular, we expect $J^{\kappa}$ to approximate the true value function for our control problem.

We will base the policy gradient (PG) step, on the representation (2.4) of the gradient, which we developed in Section 2.1. To fit a model-free setting, we consider that in general we do not know the exact form of $c$, but instead that at any point in time $s \in[t, T]$, we are able to observe our cumulative reward up to the current time,

$$
R_{s}^{t, x, a}:=\int_{t}^{s} f\left(r, X_{r}^{t, x, a}, I_{r}^{t, a}\right) d r-\int_{(t, s]} \int_{A} c\left(r, X_{r-}^{t, x, a}, I_{r-}^{t, a}, e\right) \nu(d r, d e) .
$$

This enables us, by observing our accumulated reward right before and after we change our action, so $R_{\tau_{n}-}^{t, x, a}$ before the jump and $R_{\tau_{n}}^{t, x, a}$ after the jump, to compute the cost term appearing in (2.4) as follows

$$
c\left(\tau_{n}, X_{\tau_{n}-}^{t, x, a}, I_{\tau_{n}-}^{t, a}, I_{\tau_{n}}^{t, a}\right)=R_{\tau_{n}-}^{t, x, a}-R_{\tau_{n}}^{t, x, a},
$$

and thus obtain the following formula for the policy gradient,

$$
\begin{aligned}
\nabla_{\theta} J(t, x, a, \theta)=\mathbb{E}^{\theta} & {\left[\int_{(t, T]} \int_{A} \nabla_{\theta}\left(\log \lambda^{\theta}\right)\left(e \mid s, X_{s-}^{t, x, a}, I_{s-}^{t, a}\right)\right.} \\
& \left.\cdot\left(J\left(s, X_{s}^{t, x, a}, e, \theta\right)-J\left(s, X_{s-}^{t, x, a}, I_{s-}^{t, a}, \theta\right)+R_{s}^{t, x, a}-R_{s-}^{t, x, a}\right) \nu(d s, d e)\right] .
\end{aligned}
$$

For the policy evaluation (PE) step, we can for example utilise a martingale loss function based on approach introduced in [10. Their approach is based on the observation that for the true value function $v^{\theta}:=J(\cdot, \theta)$ associated with a fixed policy $\lambda^{\theta}$, the process

$$
\left(v^{\theta}\left(s, X_{s}^{t, x, a}, I_{s}^{t, a}\right)+R_{s}^{t, x, a}\right)_{s \in[t, T]}
$$

is a martingale under $\mathbb{P}^{\theta}$, where $\left(X^{t, x, a}, I^{t, a}\right)$ follow the intensity policy $\lambda^{\theta}$. Now using that at terminal time the value function $v^{\theta}$ is just the terminal reward $g$, so

$$
v^{\theta}\left(T, X_{T}^{t, x, a}, I_{T}^{t, a}\right)+R_{T}^{t, x, a}=g\left(X_{T}^{t, x, a}, I_{T}^{t, a}\right)+R_{T}^{t, x, a},
$$

we can conclude that for all $s \in[t, T]$, under $\mathbb{P}^{\theta}$,

$$
v^{\theta}\left(s, X_{s}^{t, x, a}, I_{s}^{t, a}\right)+R_{s}^{t, x, a}=\underset{\xi \text { is } \mathcal{F}_{s}^{X, \nu} \text {-measurable }}{\arg \min } \mathbb{E}^{\theta}\left[\left|g\left(X_{T}^{t, x, a}, I_{T}^{t, a}\right)+R_{T}^{t, x, a}-\xi\right|^{2}\right] .
$$

Since $J^{\kappa}$ is intended to approximate $v^{\theta}$, this motivates us to consider the following martingale loss for our learned reward functional $J^{\kappa}$,

$$
\operatorname{ML}\left(J^{\kappa}\right):=\frac{1}{2} \mathbb{E}^{\theta}\left[\int_{(t, T]} \int_{A}\left|J^{\kappa}\left(s, X_{s}^{t, x, a}, I_{s}^{t, a}\right)+R_{s}^{t, x, a}-g\left(X_{T}^{t, x, a}, I_{T}^{t, a}\right)-R_{T}^{t, x, a}\right|^{2} \nu(d s, d e)\right] .
$$

This loss, in essence, quantifies how much we deviate from the martingale characterisation above. Another possible choice for $\operatorname{ML}\left(J^{\kappa}\right)$ would e.g. be $\mathbb{E}^{\theta}\left[\int_{t}^{T} \mid J^{\kappa}\left(s, X_{s}^{t, x, a}, I_{s}^{t, a}\right)+R_{s}^{t, x, a}-\right.$ $\left.g\left(X_{T}^{t, x, a}, I_{T}^{t, a}\right)-\left.R_{T}^{t, x, a}\right|^{2} d s\right]$, which has been considered by 10. To be able to learn the reward
functional for a policy $\theta$, we will update our estimate $\kappa$ using the martingale loss $\operatorname{ML}\left(J^{\kappa}\right)$ by computing

$$
\begin{aligned}
& \nabla_{\kappa} \operatorname{ML}\left(J^{\kappa}\right) \\
& =\mathbb{E}^{\theta}\left[\int_{(t, T]} \int_{A}\left(J^{\kappa}\left(s, X_{s}^{t, x, a}, I_{s}^{t, a}\right)+R_{s}^{t, x, a}-g\left(X_{T}^{t, x, a}, I_{T}^{t, a}\right)-R_{T}^{t, x, a}\right) \nabla_{\kappa} J^{\kappa}\left(s, X_{s}^{t, x, a}, I_{s}^{t, a}\right) \nu(d s, d e)\right] .
\end{aligned}
$$

By combining both steps, we then obtain the following generic actor-critic algorithm for randomised control problem in a model-free setting.

```
Algorithm 1: Offline-episodic actor-critic algorithm
    Input: initial state \(x_{0}\), initial action \(a_{0}\), parametrised family of reward functions
                \(\left(J^{\kappa}\right)_{\kappa}\), parametrised family of randomised intensity actions \(\left(\lambda^{\theta}\right)_{\theta}\), initial
                learning rates \(\eta_{\kappa}, \eta_{\theta}\), learning schedule \(l(\cdot)\)
    Output: learned value function \(J^{\kappa}\), optimal randomised control \(\lambda^{\theta}\)
    initialise \(\kappa, \theta\)
    for episode \(j=1, \ldots\) do
            simulate \(\left(X_{t}, I_{t}\right)_{t \in[0, T]}\) starting from \(\left(X_{0}, I_{0}\right)=\left(x_{0}, a_{0}\right)\) according to the policy \(\lambda^{\theta}\)
            and observe the accumulated running reward \(\left(R_{t}\right)_{t \in[0, T]}\) and the terminal reward
            \(G_{T}=g\left(X_{T}, I_{T}\right)\)
            compute \(\nabla_{\kappa} \operatorname{ML}\left(J^{\kappa}\right) \leftarrow \sum_{I_{t-\neq I_{t}}}\left(J^{\kappa}\left(t, X_{t}, I_{t}\right)+R_{t}-G_{T}-R_{T}\right) \nabla_{\kappa} J^{\kappa}\left(t, X_{t}, I_{t}\right)\)
            compute
            \(\nabla_{\theta} J^{\kappa} \leftarrow \sum_{I_{t-} \neq I_{t}}\left(J^{\kappa}\left(t, X_{t}, I_{t}\right)-J^{\kappa}\left(t, X_{t-}, I_{t-}\right)+R_{t}-R_{t-}\right) \nabla_{\theta}\left(\log \lambda^{\theta}\right)\left(I_{t} \mid t, X_{t-}, I_{t-}\right)\)
            update \(\kappa \leftarrow \kappa-\eta_{\kappa} l(j) \nabla_{\kappa} \operatorname{ML}\left(J^{\kappa}\right)\)
            update \(\theta \leftarrow \theta+\eta_{\theta} l(j) \nabla_{\theta} J^{\kappa}\)
    end
```

Finally, it is important to note that while this algorithm addresses the randomised problem, our primary interest lies in (non-randomised) stochastic control problems; their randomised counterparts serve as tools for handling these control problems. Specifically, our objective will in general not be to find the optimal intensity $\lambda^{\theta}$, but rather to find the optimal (non-randomised) control $\alpha$. Therefore, in the next Section 3, we will discuss in more detail how to utilise this algorithm to find the optimal control for the corresponding stochastic control problems.

## 3 Application to stochastic control problems

In this section, we consider general stochastic control problems for which a randomised formulation in the form of Section 2 exists. The classical problem is the case of controlled Markov processes $X^{\alpha}$, e.g., driven by diffusion processes, with regular controls $\alpha$ valued in $A$, and where the objective is to maximise over $\alpha$ a criterion in the form

$$
\mathrm{J}(\alpha)=\mathbb{E}\left[g\left(X_{T}^{\alpha}\right)+\int_{0}^{T} f\left(t, X_{t}^{\alpha}, \alpha_{t}\right) d t\right] .
$$

The corresponding randomised formulation is the one described in Section 2 with $g(x)$ depending only on $x, c \equiv 0$, and the key result, proved in [14], see also [5], is the statement that the two value functions coincide, namely:

$$
\sup _{\alpha} \mathrm{J}(\alpha)=\sup _{\lambda \in \mathcal{V}} \mathbb{E}^{\lambda}\left[g\left(X_{T}\right)+\int_{0}^{T} f\left(t, X_{t}, I_{t}\right) d t\right] .
$$

Such randomised formulations have been developed for a large class of continuous time control problems, including, but not limited to, impulse control problems in [13, optimal stopping in [6], or optimal switching problems in [1, 4] as it will be illustrated in the next section. The core idea behind the randomisation framework is to replace the control by a random point process, usually a Poisson point process, whose intensity becomes the new control, which results in a randomised problem in form of Section 2 The advantage of this randomised formulation is that it provides a unified framework for many different classes of control problems, and not only for continuous time problems but even includes stochastic control in discrete time on deterministic and/or random grids. Our goal is then to develop an actor-critic algorithm for the original stochastic control problem by utilising its randomised counterpart and its actor-critic algorithm derived in Section 2.2 However, the drawback is that the randomised problem is not directly equivalent to the original problem. While one can often show that the value functions of both problems coincide, it is not trivial how to recover an optimal control for the original problem just from studying the randomised formulation. In particular, it is also not clear whether the randomised formulation even has an optimal (randomised) control - in general this will not be the case.

Let us for simplicity of presentation assume that $\hat{\nu}(d s, d e)=\bar{\nu}(d s) \mu\left(d e \mid s, X_{s-}, I_{s-}\right)$, where $\bar{\nu}$ and $\mu$ are both non-random. Then $\bar{\nu}$ describes the distribution of points in $\nu$ in time, and $\mu$ is the kernel transition probability describing the mark distributions of such points. Similarly, we split $\lambda^{\theta}$ into an intensity $\Lambda^{\theta}$ for new points and a probability density $\bar{\lambda}^{\theta}$ for their marks as follows
$\Lambda^{\theta}(s, x, a):=\int_{A} \lambda^{\theta}(e \mid s, x, a) \mu(d e \mid s, x, a), \quad \bar{\lambda}^{\theta}(e \mid s, x, a):=\frac{\lambda^{\theta}(e \mid s, x, a)}{\Lambda^{\theta}(s, x, a)}, \quad(s, x, a) \in[0, T] \times \mathbb{R}^{d} \times A$.
Let us further assume that $\mu(\{a\} \mid s, x, a)>0$ for all $(s, x, a) \in[0, T] \times \mathbb{R}^{d} \times A$, so that at any point there is a non-negative probability that the jump of $\bar{\nu}$ does not induce a real jump in $I$. Given a function $\Lambda$, we now denote by $\Theta_{\leq \Lambda}$ (resp. $\Theta_{=\Lambda}$ ) the set of all $\theta \in \Theta$ such that $\Lambda^{\theta} \leq \Lambda$ (resp. $\Lambda^{\theta}=\Lambda$ ). Then we note that for every $\theta \in \Theta_{\leq \Lambda}$, the process $\lambda(e \mid s, x, a):=$ $\lambda^{\theta}(e \mid s, x, a)+\left(\Lambda(s, x, a)-\Lambda^{\theta}(s, x, a)\right) \frac{1}{\mu(\{a\} \mid s, x, a)} \in \mathcal{V}$ emulates the control $\lambda^{\theta}$ in the sense that $\mathbb{P}_{(X, I)}^{\lambda}=\mathbb{P}_{(X, I)}^{\theta}$. Therefore, supposing that $\left(\lambda_{\theta}\right)_{\theta \in \Theta}$ is sufficiently dense in $\mathcal{V}$, we see that $\lambda$ can be approximated by $\left(\lambda^{\theta_{n}}\right)_{n}$ such that $\theta_{n} \in \Theta_{=\Lambda}$, which shows that

$$
\sup _{\theta \in \Theta_{\leq \Lambda}} J(t, x, a, \theta)=\sup _{\theta \in \Theta=\Lambda} J(t, x, a, \theta) .
$$

This motivates us to view $\Lambda$ as some kind of inverse step size, since as $\Lambda \rightarrow \infty($ resp. $\Lambda \bar{\nu}(\{s\}) \uparrow 1$ if $\bar{\nu}(\{s\})>0)$, we see that

$$
\sup _{\theta \in \Theta=\Lambda} J(t, x, a, \theta) \rightarrow \sup _{\theta \in \Theta} J(t, x, a, \theta) .
$$

Thus, we introduce the following restricted parameter sets for $\theta$, for $\Lambda_{c} \geq 0$ and $0 \leq \Lambda_{d} \leq 1$,

$$
\Theta_{\Lambda_{c}, \Lambda_{d}}=\left\{\theta \in \Theta \left\lvert\, \Lambda^{\theta}(s, x, a)=\Lambda_{c} \mathbb{1}_{\{\bar{\nu}(\{s\})=0\}}+\frac{\Lambda_{d}}{\bar{\nu}(\{s\})} \mathbb{1}_{\{\bar{\nu}(\{s\})>0\}}\right., \text { for all }(s, x, a)\right\} \subseteq \Theta
$$

Then, $\sup _{\theta \in \Theta_{\Lambda_{c}, \Lambda_{d}}} J \rightarrow \sup _{\theta \in \Theta} J$ as long as $\Lambda_{c} \rightarrow \infty$ and $\Lambda_{d} \uparrow 1$. Further, since while optimising over $\Theta_{\Lambda_{c}, \Lambda_{d}}$, the intensity is fixed, it is equivalent to optimise instead over the probability densities $\left(\bar{\lambda}_{\theta}\right)_{\theta \in \Theta_{\Lambda_{c}, \Lambda_{d}}}$, which by construction satisfy $\int_{A} \bar{\lambda}^{\theta}(e \mid \cdot) \mu(d e \mid \cdot) \equiv 1$ and $\bar{\lambda}^{\theta} \geq 0$. Note that this family, if $\Theta$ is sufficiently exhaustive, does not actually depend on $\Lambda_{c}, \Lambda_{d}$ anymore. Thus, by imposing an intensity schedule $\Lambda_{c}$ and $\Lambda_{d}$ ensuring that $\Lambda_{c} \rightarrow \infty$ and $\Lambda_{d} \uparrow 1$, we obtain Algorithm 2 .

At the same time, we recall that solving the randomised problem was however not our original goal. Instead, it serves as a tool for solving the original (non-randomised) problem. In particular, the $\bar{\lambda}^{\theta}$ we obtain from Algorithm 1 is not our desired control. We recall that instead $\bar{\lambda}^{\theta}$ represents the intensity for $I$, and the process $I$ then actually plays the role of the sought-after control process $\alpha$. Therefore, let us define for each $\theta$ also a control $\alpha^{\theta}$ as e.g. the arg max of the distribution $\bar{\lambda}^{\theta}(e \mid s, x, a) \mu(d e \mid s, x, a)$. The motivation is that at each jump $s$ of $I$, we draw our new control $I_{s}$ from the distribution $\bar{\lambda}^{\theta}\left(e \mid s, X_{s-}, I_{s-}\right) \mu\left(d e \mid s, X_{s-}, I_{s-}\right)$, and if the intensity $\Lambda^{\theta} \rightarrow \infty\left(\right.$ resp. $\Lambda^{\theta} \bar{\nu}(\{s\}) \uparrow 1$ if $\left.\bar{\nu}(\{s\})>0\right)$, then we are essentially able to draw a new control at every time point $s \in[t, T]$, just as in the original control problem. Consequently, letting in our case $\Lambda_{c} \rightarrow \infty$ and $\Lambda_{d} \uparrow 1$, this then leads to the convergence of $\alpha^{\theta} \rightarrow \alpha_{*}$.

```
Algorithm 2: Offline-episodic actor-critic algorithm with random grids and intensity schedule
    Input: initial state \(x_{0}\), initial action \(a_{0}\), terminal time \(T\), parametrised family of
                    reward functions \(\left(J^{\kappa}\right)_{\kappa}\), baseline random grid sampling distribution \(\bar{\nu}\), baseline
                    action distribution kernel \(\mu\), parametrised family of action densities \(\left(\bar{\lambda}^{\theta}\right)_{\theta}\),
                    intensity schedule \(\Lambda_{c}(\cdot), \Lambda_{d}(\cdot)\), initial learning rates \(\eta_{\kappa}, \eta_{\theta}\), learning schedule \(l(\cdot)\)
    Output: learned value function \(J^{\kappa}\), optimal randomised control \(\lambda^{\theta}\), optimal
                (non-randomised) control \(\alpha^{\theta}\)
    initialise \(\kappa, \theta\)
    for episode \(j=1, \ldots\) do
        initialise \(\tau_{0} \leftarrow 0, r_{0} \leftarrow 0\)
        simulate point process \(U\) on \((0, T]\) with stochastic intensity
            \(\Lambda_{c}(j) \mathbb{1}_{\{\bar{\nu}(\{s\})=0\}} \bar{\nu}(d s)+\Lambda_{d}(j) \mathbb{1}_{\{\bar{\nu}(\{s\})>0\}} \delta_{s}(d s) \rightarrow\) obtain grid points \(\left(\tau_{n}\right)_{n=1, \ldots, N}\)
            for \(n=1, \ldots, N\) do
            simulate \(X_{\left[\tau_{n-1}, \tau_{n}\right)}\) from \(X_{\tau_{n-1}}=x_{n-1}\) with control \(a_{n-1}\)
            observe the new state \(x_{n-} \leftarrow X_{\tau_{n}-}\) and accumulated running reward
                \(r_{n-} \leftarrow R_{\tau_{n}-}\) at time \(\tau_{n}-\)
                simulate and update the new control \(a_{n} \sim \bar{\lambda}^{\theta}\left(e \mid \tau_{n}, x_{n-}, a_{n-1}\right) \mu\left(d e \mid \tau_{n}, x_{n-}, a_{n-1}\right)\)
                observe the new state \(x_{n} \leftarrow X_{\tau_{n}}\) and accumulated running reward \(r_{n} \leftarrow R_{\tau_{n}}\)
                after updating the control at time \(\tau_{n}\)
            end
            simulate \(X_{\left[\tau_{N}, T\right]}\) with control \(a_{N}\)
            observe the final state \(x_{N+1} \leftarrow X_{T}\) and set \(a_{N+1} \leftarrow a_{N}, \tau_{N+1} \leftarrow T\)
            observe the final accumulated running reward \(r_{N+1} \leftarrow R_{T}\) and the terminal reward
            \(G_{T}=g\left(X_{T}, I_{T}\right)\) at time \(T\)
            compute \(\nabla_{\kappa} \operatorname{ML}\left(J^{\kappa}\right) \leftarrow \sum_{a_{n} \neq a_{n-1}}\left(r_{n}+J^{\kappa}\left(\tau_{n}, x_{n}, a_{n}\right)-G_{T}-r_{N+1}\right) \nabla_{\kappa} J^{\kappa}\left(\tau_{n}, x_{n}, a_{n}\right)\)
            compute \(\nabla_{\theta} J^{\kappa} \leftarrow\)
            \(\sum_{a_{n} \neq a_{n-1}}\left(J^{\kappa}\left(\tau_{n}, x_{n}, a_{n}\right)-J^{\kappa}\left(\tau_{n}, x_{n-}, a_{n-1}\right)+r_{n}-r_{n-}\right) \nabla_{\theta}\left(\log \bar{\lambda}^{\theta}\right)\left(a_{n} \mid \tau_{n}, x_{n-}, a_{n-1}\right)\)
            update \(\kappa \leftarrow \kappa-\eta_{\kappa} l(j) \nabla_{\kappa} \operatorname{ML}\left(J^{\kappa}\right)\)
            update \(\theta \leftarrow \theta+\eta_{\theta} l(j) \nabla_{\theta} J^{\kappa}\)
    end
    obtain \(\alpha^{\theta}(t, x, a)\) as the \(\arg\) max of the probability distribution \(\bar{\lambda}^{\theta}(e \mid t, x, a) \mu(d e \mid t, x, a)\)
```

Finally, we want to conclude this section by giving a version of the actor-critic algorithm but with a "fixed step size". This is the version we will also use later in Section 4. So we will fix the intensity $\Lambda$ and thus the base process $\tilde{\nu}:=\Lambda \bar{\nu}$, and now optimise again over the parametrised family of action densities $\left(\bar{\lambda}^{\theta}\right)_{\theta}$, which results in the following Algorithm 3 .

Remark 3.1. This leads to a flexible framework accommodating various types of sampling grids, such as

- deterministic discrete grids, by choosing $\tilde{\nu}=\sum_{k=0}^{N} \delta_{\frac{k T}{N}}(d s)$, for $N \geq 2$,
- random discrete grids, by setting $\tilde{\nu}=\sum_{k=0}^{N} p_{\text {samp }} \delta_{\frac{k T}{N}}(d s)$, for $N \geq 2$ and $p_{\text {samp }} \in(0,1]$,
- Poisson grids are possible with $\tilde{\nu}=\lambda d$ for some intensity $\lambda>0$.

In Section 4, we will compare the choice of deterministic and random discrete grids through two numerical examples of optimal switching problems.

```
Algorithm 3: Offline-episodic actor-critic algorithm with random grids
    Input: initial state \(x_{0}\), initial action \(a_{0}\), terminal time \(T\), parametrised family of
                    reward functions \(\left(J^{\kappa}\right)_{\kappa}\), random grid sampling distribution \(\tilde{\nu}\), baseline action
                    distribution kernel \(\mu\), parametrised family of action densities \(\left(\bar{\lambda}^{\theta}\right)_{\theta}\), initial
            learning rates \(\eta_{\kappa}, \eta_{\theta}\), learning schedule \(l(\cdot)\)
    Output: learned value function \(J^{\kappa}\), optimal randomised control \(\lambda^{\theta}\), optimal
                (non-randomised) control \(\alpha^{\theta}\)
    initialise \(\kappa, \theta\)
    for episode \(j=1, \ldots\) do
        initialise \(\tau_{0} \leftarrow 0, r_{0} \leftarrow 0\)
        simulate point process on \((0, T]\) with stochastic intensity \(\tilde{\nu} \rightarrow\) obtain grid points
        \(\left(\tau_{n}\right)_{n=1, \ldots, N}\)
        for \(n=1, \ldots, N\) do
            simulate \(X_{\left[\tau_{n-1}, \tau_{n}\right)}\) from \(X_{\tau_{n-1}}=x_{n-1}\) with control \(a_{n-1}\)
            observe the new state \(x_{n-} \leftarrow X_{\tau_{n}-}\) and accumulated running reward
                \(r_{n-} \leftarrow R_{\tau_{n}-}\) at time \(\tau_{n}-\)
            simulate and update the new control \(a_{n} \sim \bar{\lambda}^{\theta}\left(e \mid \tau_{n}, x_{n-}, a_{n-1}\right) \mu\left(d e \mid \tau_{n}, x_{n-}, a_{n-1}\right)\)
            observe the new state \(x_{n} \leftarrow X_{\tau_{n}}\) and accumulated running reward \(r_{n} \leftarrow R_{\tau_{n}}\)
                after updating the control at time \(\tau_{n}\)
            end
            simulate \(X_{\left[\tau_{N}, T\right]}\) with control \(a_{N}\)
            observe the final state \(x_{N+1} \leftarrow X_{T}\) and set \(a_{N+1} \leftarrow a_{N}, \tau_{N+1} \leftarrow T\)
            observe the final accumulated running reward \(r_{N+1} \leftarrow R_{T}\) and the terminal reward
                \(G_{T}=g\left(X_{T}, I_{T}\right)\) at time \(T\)
            compute \(\nabla_{\kappa} \operatorname{ML}\left(J^{\kappa}\right) \leftarrow \sum_{a_{n} \neq a_{n-1}}\left(r_{n}+J^{\kappa}\left(\tau_{n}, x_{n}, a_{n}\right)-G_{T}-r_{N+1}\right) \nabla_{\kappa} J^{\kappa}\left(\tau_{n}, x_{n}, a_{n}\right)\)
            compute \(\nabla_{\theta} J^{\kappa} \leftarrow\)
            \(\sum_{a_{n} \neq a_{n-1}}\left(J^{\kappa}\left(\tau_{n}, x_{n}, a_{n}\right)-J^{\kappa}\left(\tau_{n}, x_{n-}, a_{n-1}\right)+r_{n}-r_{n-}\right) \nabla_{\theta}\left(\log \bar{\lambda}^{\theta}\right)\left(a_{n} \mid \tau_{n}, x_{n-}, a_{n-1}\right)\)
            update \(\kappa \leftarrow \kappa-\eta_{\kappa} l(j) \nabla_{\kappa} \operatorname{ML}\left(J^{\kappa}\right)\)
            update \(\theta \leftarrow \theta+\eta_{\theta} l(j) \nabla_{\theta} J^{\kappa}\)
    end
    define \(\alpha^{\theta}(t, x, a)\) as the \(\arg\) max of the probability distribution \(\bar{\lambda}^{\theta}(e \mid t, x, a) \mu(d e \mid t, x, a)\)
```


## 4 Numerical experiments for switching problems using neural networks

### 4.1 Optimal switching problem

In this paragraph, we recall the connection between optimal switching problems with their randomised formulation following [1]. Note that this randomisation method has been extended
to large class of further problems including impulse control, optimal stopping and regular control problems, and thus such problems also fit into the unified setting for our policy gradient method.

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a complete probability space carrying an $m$-dimensional Brownian motion $W$ and let $\mathbb{F}^{W}$ be its generated filtration. Let $A=\{1, \ldots, N\}$ for some fixed $N \in \mathbb{N}$ denote the finite action set. A switching control is an $A$-valued piece-wise constant process of the form

$$
\alpha=a \mathbb{1}_{\left[t, \tau_{0}\right)}+\sum_{n \in \mathbb{N}} \xi_{n} \mathbb{1}_{\left[\tau_{n}, \tau_{n+1}\right)},
$$

where $\left(\tau_{n}\right)_{n}$ is an increasing sequence of stopping times such that $\tau_{n} \rightarrow \infty \mathbb{P}$-a.s., $\left(\xi_{n}\right)_{n}$ is a sequence of $A$-valued random variables such that $\xi_{n}$ is $\mathcal{F}_{\tau_{n}}^{W}$-measurable and $a \in A$ is a fixed initial control. Given an initial value $x \in \mathbb{R}^{d}$, we further consider the controlled state dynamics as the solution to the stochastic differential equation

$$
\begin{equation*}
X_{s}^{t, x, \alpha}=x+\int_{t}^{s} b\left(r, X_{r}^{t, x, \alpha}, \alpha_{r}\right) d r+\int_{t}^{s} \sigma\left(r, X_{r}^{t, x, \alpha}, \alpha_{r}\right) d W_{r}+\sum_{t<\tau_{n} \leq s} \gamma\left(\tau_{n}, X_{\tau_{n}-}^{t, x, \alpha}, \alpha_{\tau_{n}-}, \alpha_{\tau_{n}}\right) . \tag{4.1}
\end{equation*}
$$

We denote by $\mathcal{A}$ as the set of as all switching controls $\alpha$. Note that together with the following Assumption A, this ensures that the above state dynamics (4.1) are well-defined.

Assumption A. The coefficients $b, \sigma$ and $\gamma$ are Lipschitz and of linear growth w.r.t. x: there exists a positive constant $C$ s.t. for all $t \in[0, T], x, x^{\prime} \in \mathbb{R}^{d}$, a, $a^{\prime} \in A$,

$$
\begin{aligned}
&\left|b(t, x, a)-b\left(t, x^{\prime}, a\right)\right|+\left|\sigma(t, x, a)-\sigma\left(t, x^{\prime}, a\right)\right|+\left|\gamma\left(t, x, a, a^{\prime}\right)-\gamma\left(t, x^{\prime}, a, a^{\prime}\right)\right| \leq C\left|x-x^{\prime}\right| \\
&|b(t, x, a)|+|\sigma(t, x, a)|+\left|\gamma\left(t, x, a, a^{\prime}\right)\right| \leq C(1+|x|) .
\end{aligned}
$$

Our goal is now to maximise the following reward functional

$$
\mathrm{J}(t, x, a, \alpha):=\mathbb{E}\left[g\left(X_{T}^{t, x, \alpha}, \alpha_{T}\right)+\int_{t}^{T} f\left(s, X_{s}^{t, x, \alpha}, \alpha_{s}\right) d s-\sum_{t<\tau_{n} \leq T} c\left(\tau_{n}, X_{\tau_{n}-}^{t, x, \alpha}, \alpha_{\tau_{n}-}, \alpha_{\tau_{n}}\right)\right],
$$

and we define the value function as follows

$$
V(t, x, a):=\sup _{\alpha \in \mathcal{A}} \mathrm{J}(t, x, a, \alpha) .
$$

We make the standard assumptions on the gain and cost functions:
Assumption B. The reward functions $f, g$ and the cost function are continuous w.r.t. the $x$ argument with quadratic growth condition: there exists some positive constant $C$ s.t. for all $t \in$ $[0, T], x \in \mathbb{R}^{d}, a, a^{\prime} \in A$,

$$
|f(t, x, a)|+|g(x, a)|+\left|c\left(t, x, a, a^{\prime}\right)\right| \leq C\left(1+|x|^{2}\right) .
$$

To formulate the randomised version of this problem as in Section 2, we introduce an independent random point process $\nu$ on $[0, T] \times A$ with predictable projection $\hat{\nu}$. Correspondingly,
we define the $A$-valued process

$$
I_{s}^{t, a}=a+\int_{(t, s]} \int_{A}\left(e-I_{r-}^{t, a}\right) \nu(d r, d e), \quad s \in[t, T],
$$

which will replace our control process. In particular, the state process will follow the following uncontrolled state dynamics
$X_{s}^{t, x, a}=x+\int_{t}^{s} b\left(r, X_{r}^{t, x, a}, I_{r}^{t, a}\right) d r+\int_{t}^{s} \sigma\left(r, X_{r}^{t, x, a}, I_{r}^{t, a}\right) d W_{r}+\int_{(t, s]} \int_{A} \gamma\left(r, X_{r-}^{t, x, a}, I_{r-}^{t, a}, e\right) \nu(d r, d e)$.
Our set of control will instead now be the set $\mathcal{V}$ of $\mathbb{F}^{W, \nu} \otimes \mathcal{B}(A)$-predictable, essentially bounded processes $\lambda$ such that there exists a with respect to $\mathbb{P}$ absolutely continuous probability measure $\mathbb{P}^{\lambda} \ll \mathbb{P}$ under which $\nu$ is a random point process with predictable projection $\lambda_{s}(e) \hat{\nu}(d s, d e)$, see also (2.3) for a characterisation of such $\lambda \in \mathcal{V}$. Then the reward functional is defined by

$$
J(t, x, a, \lambda):=\mathbb{E}^{\mathbb{P}^{\lambda}}\left[g\left(X_{T}^{t, x, a}, I_{T}^{t, a}\right)+\int_{t}^{T} f\left(s, X_{s}^{t, x, a}, I_{s}^{t, a}\right) d s-\int_{(t, T]} c\left(s, X_{s-}^{t, x, a}, I_{s-}^{t, a}, e\right) \nu(d s, d e)\right]
$$

and we introduce the following randomised value function

$$
V^{\mathcal{R}}(t, x, a):=\sup _{\lambda \in \mathcal{V}} J(t, x, a, \lambda) .
$$

Bouchard [1] studied the case where $\nu$ is a Poisson point process with compensator $\hat{\nu}(d s, d e)=$ $\sum_{a \in A} \delta_{a}(d e) d s$, for which the set of admissible randomised controls $\mathcal{V}$ then reduces to all $\mathbb{F}^{W, \nu} \otimes$ $\mathcal{B}(A)$-predictable, essentially bounded processes $\lambda$. Under some additional regularity and growth assumptions, it is proved the following equivalence result between the optimal switching and the randomised problem.

Theorem 4.1 ( 1 , Theorem 2.1]). Let Assumptions A and $B$ hold and $\nu$ have a predictable projection of the form $\hat{\nu}(d s, d e)=\sum_{a \in A} \delta_{a}(d e) d s$. Further assume that the regularity assumptions [1. Assumptions H1, H2] are satisfied. Then value functions of both problems coincide, that is $V=V^{\mathcal{R}}$.

Remark 4.2. While usually $\nu$ is chosen as Poisson point process, any sufficiently dense point measure, under suitable assumptions, would work for such an equivalence result. In particular, starting from any given point process, even with deterministic atoms as in the case of a deterministic grid, and by e.g. adding additional points sampled from a Poisson point process, one would obtain such a sufficiently dense point measure.

In the sequel for our numerical experiments, we consider an optimal switching problem where part of the state which is not controlled is continuous, while the other part is controlled and takes discrete values. Therefore the randomised controlled state is modeled by a discrete Markov chain described by probability transitions which are functions of the global state. At convergence, we expect that these probabilities converge either to 1 or 0 , are discontinuous in time for a fixed state, so that they cannot be represented as functions of time using neural networks. Consequently, we use a deterministic uniform grid of $N$ dates on $[0, T]$. We note $t_{n}=n \Delta t$ where $\Delta t=\frac{T}{N}$ with $\bar{N}=N-1$ the number of time steps. At each time step and each possible
state, a neural network is used to describe the transition probabilities from one state to the other on.

We then propose :

- either to sample time randomly on the deterministic grid : the number of points $\tilde{N}$ chosen on the grid is sampled using a binomial distribution with a probability $p_{\text {samp }}$ and a number of trials equal to $N-2$. Then the points from the grid are chosen randomly with an uniform law,
- or to take the $N$ points grid corresponding to $p_{\text {samp }}=1$.

We denote by $\left(\tau_{k}\right)_{k \geq 0}$ the random lattice sampled from the deterministic lattice with $\tau_{0}=0$, and complete it with the convention that $\tau_{p}=T$ for $k<p \leq N$ if $\tau_{k}=T$. We set $[k]=\frac{\tau_{k}}{\Delta t}$ the random grid index associated with values in $[0, N]$.

Remark 4.3. The fact that the control has to be modeled at each time step by different networks to get good results for degenerated controlled states with constraints was already shown in the case of reservoir optimization in 18 .

Next, depending on the problem, it may be interesting to take a representation of the reward functional $J$ different form the one proposed by $\mathrm{ML}_{\tau}\left(J^{\kappa}\right)$ and in the different examples below we detail the representation taken.

In the two examples below, we model the energy curve using the classical HJM model as [18]:

$$
\begin{equation*}
\frac{d F(t, T)}{F(t, T)}=e^{-\beta(T-t)} \sigma d W_{t} \tag{4.2}
\end{equation*}
$$

where $W_{t}$ is a one-dimensional Brownian motion defined on a probability space $(\Omega, \mathcal{F}, \mathbb{P})$. The spot price is then equal to $S_{t}=F(t, t)$. As numerical example, we take $T=30, F(0, t)=$ $90+10 \cos \left(2 \pi\left\lfloor\frac{t}{30}\right\rfloor\right), \beta=0.15, \sigma=0.5$.

### 4.2 Starting and stopping in physical assets

We consider the problem of a thermal asset generating power as in $[8], \sqrt{16}$. The asset has a production cost of $K$ per time unit, and has two states : either on (state 1 ) or off (state 0 ) and the switching control $\alpha=\left(\alpha_{t}\right)_{t}$ is

$$
\alpha_{t}=\alpha_{0} \mathbb{1}_{\left[\tau_{0}, \tau_{1}\right)}+\sum_{n>0} \xi_{n} \mathbb{1}_{\left[\tau_{n}, \tau_{n+1}\right)}(t), \quad 0<t \leq T
$$

with $\alpha_{0}=1$ (the asset is on at $t=0$ ), and the random variables $\left(\xi_{n}\right)_{n}$ denote the sequence of operating regimes valued in $A=\{0,1\}$, representing the decisions to stop or run the production. There is a fixed cost for switching from a mode to another one, namely $c_{0,1}$ (resp. $c_{1,0}$ ) for starting (resp. stopping) the production.

The manager of the power asset aims to maximize over $\alpha$ the expected global profit:

$$
\mathrm{J}(\alpha)=\mathbb{E}\left[\int_{0}^{T} f\left(S_{t}, \alpha_{t}\right) d t-\sum_{n} c_{\alpha_{\tau_{n}}, \alpha_{\tau_{n+1}}}\right]
$$

with running profit functions $f(s, 0)=0$, and $f(s, 1)=s-K$.

- As for the control, we model the switching probability at each time step $n$ by two neural networks depending on the price, and using a sigmoid function at the output layer, $\bar{\lambda}^{\theta_{n, i}}(S)$ takes values in $[0,1]$ for $i=0,1$ with parameters $\theta_{n, i}$ and such that for $i \in\{0,1\}, \bar{\lambda}^{\theta_{n, i}}$ is the probability that, given the state $i$ in $t_{n}^{-}$, the asset will change to state $1-i$ in $t_{n}$. Here, $\theta=\left(\left(\theta_{n, i}\right)_{n=1, \bar{N}-1}\right)_{i=0,1}$.
- As for the value function $J$ we use similarly for each time step $n$ two neural networks depending on the price $S:\left(J^{\kappa_{n, i}}(S)\right)_{i=0,1}$ taking values in $\mathbb{R}$ with parameters $\kappa_{n, i}$ where $J^{\kappa_{n, i}}$ is the value function in state $i$ for $i \in\{0,1\}$. Here, $\kappa=\left(\left(\kappa_{n, i}\right)_{n=0, \bar{N}-1}\right)_{i=0,1}$. We also take the convention $J^{\kappa_{[N], i}}=0$ for $i=0,1$.

For this example, to estimate the reward function $J$, we propose to minimize the loss function

$$
\sum_{k=0}^{\bar{N}-1} \mathbb{E}\left[\left|J^{\kappa_{[k], \xi_{k}}}\left(S_{\tau_{k}}\right)-\sum_{n=0}^{\bar{N}-1}\left(S_{t_{n}}^{\tau_{k}}-K\right) \Delta t \mathbb{1}_{\left\{\alpha_{t_{n}}^{\xi_{k}}=1\right\}} \mathbb{1}_{\left\{t_{n} \geq \tau_{k}\right\}}+\sum_{n \geq k} c_{\alpha_{\tau_{n}}^{\xi_{k}, \alpha_{\tau_{n+1}}^{\xi_{k}}}} \mathbb{1}_{\left\{\tau_{n+1}<T\right\}}\right|^{2}\right]
$$

with the convention $c_{i, i}=0, i=0,1$ and where for $k=0, \ldots, \bar{N}-1, \xi_{k}$ is the state regime with values in $\{0,1\}$ which is sampled uniformly. $\alpha_{t}^{\xi_{k}}$ for $t \geq \tau_{k}$ denote the regimes sampled randomly using probabilities $\left(\bar{\lambda}^{\theta}{ }^{[l], \alpha_{\tau_{l-1}}^{\xi_{k}}}\left(S_{\tau_{l}}\right)\right)_{l>k}$ starting at date $\tau_{k}$ with a value $\xi_{k}$, while $S_{t_{n}}^{\tau_{k}}$ is the asset value in $t_{n}$ obtained by sampling from its initial distribution in $\tau_{k}$ according to the asset law.

The gradient function $\nabla_{\theta} J^{\kappa}$ is estimated locally for each time step and the sum of the local gradients $D W(\theta)$ is used

$$
D W(\theta)=\sum_{n=0}^{\bar{N}-2} \mathbb{E}\left[\nabla_{\theta} \log \left(\bar{\lambda}^{\theta_{n+1, \xi_{n}}}\left(S_{t_{n+1}}\right)\right)\left(J^{\kappa_{n+1, \alpha_{t_{n+1}}^{\xi_{n}}}\left(S_{t_{n+1}}\right)-J^{\kappa_{n+1}, \xi_{n}}}\left(S_{t_{n+1}}\right)-c_{\xi_{n}, \alpha_{t_{n+1}}^{\xi_{n}}}\right)\right]
$$

where once again for each $n$ in the loop $S_{t_{n+1}}$ is sampled according the asset law at date $t_{n+1}, \xi_{k}$ is sampled uniformly in $\{0,1\}$, and $\alpha_{t_{n+1}}^{\xi_{n}}$ is sampled according the probability $\bar{\lambda}^{\theta_{n+1, \xi_{n}}}\left(S_{t_{n+1}}\right)$.

Remark 4.4. Instead of simulating the process $X$ in forward direction as in Algorithm 3 and evaluating $\nabla_{\theta} J^{\kappa}$, we use a local version of the gradient which randomly samples the state at every possible time step. This extra randomisation of the state allows us to be sure that all states are explored and gives better results. This kind of extra randomisation is generally used in classical actor critic methods, where the control is taken as a normal law with decreasing variance with iterations (see for example [15]).

We use the ADAM algorithm with $\eta_{\theta}=0.00015, \eta_{\kappa}=0.03$. As noticed in [15], it is crucial to have $\eta_{\theta} \ll \eta_{\kappa}$ to have good convergence. The tanh activation function is taken for the activation functions in the hidden layers, while the sigmoid activation function is used for the output layer for the probabilities. The batch size is equal to 10000 . The references are calculated using dynamic programming in the StOpt library [7] where regression are calculated using adapted linear regression per mesh [2] and 30 time steps so $N=31$ are used. In the deterministic case, we get a value of 86 , while using $\sigma=0.15$, the value function is equal to 146.9 . On Figure 1 , we plot the convergence of the actor critic algorithm in the stochastic case using $p_{\text {samp }}=1$. The convergence to the correct value is achieved after more than 10000 gradient iterations. In the graph, the "Function value" is obtained using the $J^{\kappa}$ approximation of $J$, while the "Gain
expectation" is obtained using the gain estimate in the simulation (the controls and grids are sampled).


Figure 1: Convergence of the actor critic algorithm for the thermal switching problem depending on the gradient iteration.

On Figure 2, we explore the effect of extra temporal randomization. Not surprisingly, using a fixed lattice to sample from degrades the results as the sampling ratio $p_{\text {samp }}$ is lower. Similarly, by fixing the sampling ratio, the results improve as we increase the number of time steps of the lattice.


Figure 2: Convergence of the actor critic algorithm (function value) for the thermal switching problem depending on the gradient iteration letting $p_{\text {samp }}$ or $N$ vary.

It is also possible to use the estimated probability and in the simulation, using an Eulerian scheme, the control is selected as the most probable. In the stochastic case, we get a value equal to 146.0 using $p_{\text {samp }}=1$ at the end of the iterations, while the value obtained with $N=61$, $p_{\text {samp }}=0.64$ is 139.1 .

We give the control obtained in simulation (taken as the one with the highest probability) using $p_{\text {samp }}=1$ on Figure 3.

$t=15$, OFF state (0)


$$
t=23, \mathrm{OFF} \text { state }(0)
$$


$t=15, \mathrm{ON}$ state (1)

$t=23, \mathrm{ON}$ state $(1)$

Figure 3: Control obtained in the stochastic thermal test case depending the state for two date.

### 4.3 A storage model

We consider the example of a battery storage valuation formulated as an optimal switching problem, see [3], [18]: the manager of the battery aims to price its real options value by optimizing over a finite horizon the dynamic decisions to inject or withdraw power. The inventory process of the battery is denoted by $\left(K_{t}\right)_{t}$, and is controlled by a switching control $\alpha=\left(\alpha_{t}\right)_{t}$ :

$$
\alpha_{t}=\alpha_{0} \mathbb{1}_{\left[\tau_{0}, \tau_{1}\right)}+\sum_{n>0} \xi_{n} \mathbb{1}_{\left[\tau_{n}, \tau_{n+1}\right)}(t), \quad 0<t \leq T
$$

where the random variables $\left(\xi_{n}\right)_{n}$ denote the sequence of operating regimes valued in $A=$ $\{-1,0,1\}$, representing the decisions to withdraw, do nothing, or inject power. At $t=0$, the battery is withdrawing power so that $\alpha_{0}=-1$.

The effort of moving from regime $i \in A$ to another regime $j \in A$ incurs a cost $c_{i, j}$ with $c_{i, i}$ $=0, c_{i, j}>0$ for $i \neq j$. The inventory is given by: $K_{t}=\int_{0}^{t} \alpha_{s} d s$, while satisfying the physical constraint: $K_{t} \in\left[0, K_{\max }\right]$, for all $t \in[0, T]$. Therefore the number of discrete states is $k_{\max }+1$ where $k_{\max }=\frac{K_{\max }}{\Delta t}$. The exogenous price of the electricity is governed by equation (4.2). The
objective of the manager is to maximize over switching control $\alpha$ the reward functional

$$
J(\alpha)=\mathbb{E}\left[\int_{0}^{T} f\left(S_{t}, \alpha_{t}\right) d t-\sum_{n} c_{\alpha_{\tau_{n}}, \alpha_{\tau_{n+1}}}\right],
$$

with a running profit function

$$
f(s, a)= \begin{cases}-s, & \text { for } a=1 \\ 0, & \text { for } a=0 \\ s, & \text { for } a=-1\end{cases}
$$

Similarly as in the previous example:

- We model the switching probability at each time step $n$ and at each inventory $k$ by a neural network $\bar{\lambda}^{\theta_{n, k}}\left(S_{t_{i}}\right)$ with parameters $\theta_{n, k}$ and with an output in $[0,1]^{9}$. When injection, do nothing and withdraw are allowed (so when $k \in\left\{1, \ldots, k_{\max }-1\right\}$ ), for $(l, m) \in\{-1,0,1\} \times$ $\{-1,0,1\}, \bar{\lambda}_{l, m}^{\theta_{n, k}}\left(S_{t_{n}}\right)$ represents the probability to go from state $l$ at $t_{n}^{-}$to state $m$ at $t_{n}$. When withdraw (respectively injection) is not allowed therefore when $k=0$ (respectively $\left.k=k_{\text {max }}\right), \bar{\lambda}_{l,-1}^{\theta_{n, k}}$ (respectively $\left.\bar{\lambda}_{l, 1}^{\theta_{n, k}}\right)$ is set to 0 . These neural networks use a softmax activation function at the output layer to satisfy that probabilities are positive with a sum equal to 1 .
- As for the value function $J$ we use similarly for each time step $n$ and for a level $k$ a neural network $J^{\kappa_{n, k}}\left(S_{t_{n}}\right)$ with parameters $\kappa_{n, k}$ and an output in dimension 3 where $J_{i}^{\kappa_{n, k}}$ is at date $t_{n}$ and level $k$, the value function in state $i$ for $i \in\{-1,0,1\}$. As previously we take the convention $J_{i}^{k_{[N], k}}=0$ for $i \in\{-1,0,1\}$ and each level $k$.

Comparing to the thermal switching asset,

- the state encompass the inventory level and we minimize the following loss function to estimate $J$ :

$$
\sum_{n=0}^{\bar{N}-1} \mathbb{E}\left[\left|J_{\alpha_{\tau_{n}}^{n}}^{\kappa_{[n]} K_{\tau_{n}}^{n}}\left(S_{\tau_{n}}^{n}\right)-\sum_{p=0}^{\bar{N}-1} f\left(S_{t_{p}}^{n}, \alpha_{t_{p}}^{n}\right) \mathbb{1}_{\left\{t_{p} \geq \tau_{n}\right\}} \Delta t+\sum_{p \geq n} c_{\alpha_{\tau_{p}}^{n}, \alpha_{\tau_{p+1}}^{n}} \mathbb{1}_{\left\{\tau_{p+1}<T\right\}}\right|^{2}\right]
$$

where at date $\tau_{n}$ in the outer summation $S_{\tau_{n}}^{n}$ is sampled according the asset law at date $\tau_{n}$, while the inventory level $K_{\tau_{n}}^{n}$ and control applied $\alpha_{\tau_{n}}^{n}$ are sampled uniformly. $S_{t_{p}}^{n}$ is the asset value at date $t_{p}>\tau_{n}$ conditionally to its value at date $\tau_{n}$, and $\alpha_{t_{p}}^{n}$ the applied control at date $t_{p}$ starting from $\alpha_{\tau_{n}}^{n}$ at date $\tau_{n}$ and obtained sampling the switching probabilities as for the thermal asset. Therefore the flow equation for the inventory level is given for $t_{p}>\tau_{n}$ by

$$
\begin{equation*}
K_{t_{p+1}}^{n}=0 \vee\left(K_{t_{p}}+\alpha_{t_{p}}^{n} \Delta t\right) \wedge K_{\text {max }} \tag{4.3}
\end{equation*}
$$

- The gradient function $D W$ is estimated with similar notations as

$$
D W(\theta)=\sum_{n=1}^{\bar{N}-1} \mathbb{E}\left[\nabla_{\theta} \log \left(\bar{\lambda}_{\alpha_{t_{n-1}}^{n-1}, \alpha_{t_{n}}^{n-1}}^{\theta_{n, K_{t_{n}}}^{n-1}}\left(S_{t_{n}}\right)\right)\left(J_{\alpha_{t_{n}}^{n-1}}^{\kappa_{n, K_{t_{n}}}^{n-1}}\left(S_{t_{n}}\right)-J_{\alpha_{t_{n-1}}^{n-1}}^{\kappa_{n, K_{t}}}\left(S_{t_{n}}\right)-c_{\alpha_{t_{n-1}, 1}^{n-1}, \alpha_{t_{n}}^{n-1}}\right)\right]
$$

where once again for each $n$ in the loop, $S_{t_{n}}$ is sampled according the asset law at date $t_{n}$, while the inventory level $K_{t_{n}}$ at date $t_{n}$ and the control $\alpha_{t_{n-1}}^{n-1}$ at $t_{n}^{-}$are sampled uniformly. The control $\alpha_{t_{n}}^{n-1}$ is sampled from the state $\left(t_{n}, K_{t_{n}}, \alpha_{t_{n-1}}^{n-1}\right)$ using the probabilities $\bar{\lambda}_{\alpha_{t_{n-1},}^{n-1}, \alpha_{t_{n}}^{n-1}}^{\theta_{n, K_{t_{n}}}^{n}}\left(S_{t_{n}}\right)$.
Remark 4.5. We have to clip values in the flow equation (4.3). A possible control for a single time step may be not admissible if $\tau_{n+1}-\tau_{n}>1$. Then if we inject (control $\alpha=1$ ), and if the control is admissible during one time step and not for two, the control is changed to 0 on the second time step and a switching cost is added. A similar adaptation is carried out in the withdrawal regime.

As numerical example, we take the following switching costs: $c_{-1,1}=c_{1,-1}=5$, and for $(i, j)$ not in $\{(-1,1),(1-1),(-1,-1),(0,0),(1,1)\}, c_{i, j}=3$. We take $K_{0}=2, K_{\max }=2$ and the reference calculated with the StOpt library using 30 time steps is 264.3 .

We use a batch size of $10000, \eta_{\theta}=0.00015$ and $\eta_{\kappa}=0.05$ with ADAM optimizers.


Figure 4: Convergence of the actor critic algorithm for the battery storage case.

On Figure 4 we observe that taking $p_{\text {samp }}=1$ allows to recover almost the exact solution and, as for the thermal asset, the results deteriorate as $p_{\text {samp }}$ decreases.

On Figures 5 and 6, we give an example of the control obtained in each regime in simulation. At each date the control with the highest probability is taken in simulation.


Figure 5: Control obtained at date 12 in the battery test case depending on the regime.


Figure 6: Control obtained at date 25 in the battery test case depending on the regime.

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[^1]:    ${ }^{1}$ Since every $\lambda$ with $\mathbb{P}^{\lambda} \ll \mathbb{P}$ can be approximated by $\left(\lambda^{n}\right)_{n} \subseteq \mathcal{V}$, this additional assumption also does not change the value function. Similar arguments are standard for randomised control problems.

